

Multivariate Functional Data Analysis

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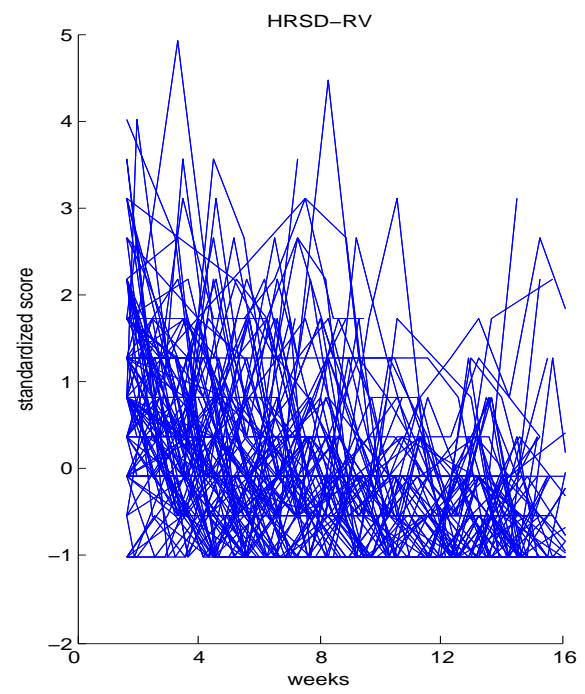
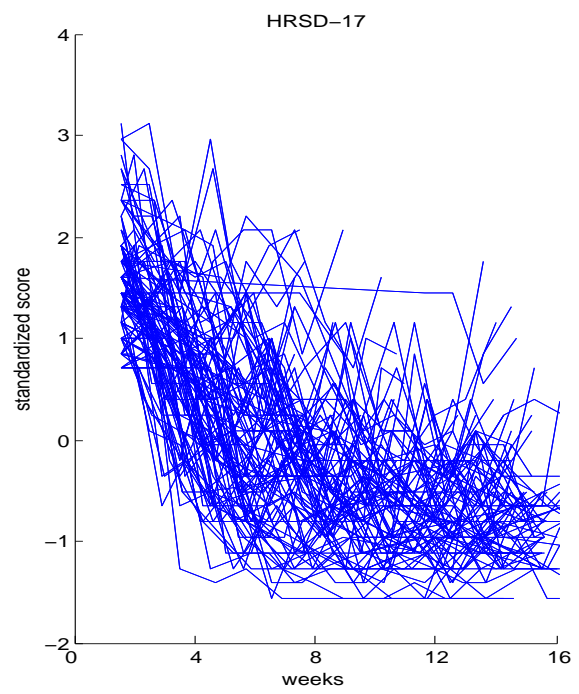
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Outline

- Motivating example
- The model
- Estimation
- Example

Motivating Example

- Psychiatric study to compare psychotherapy to pharmacotherapy at the University of Pisa, Italy. Total of 103 acutely depressed subjects.
- **Response variables:** two subscales of the clinician-administered Hamilton Depression rating Scale.
 - HRSD-17 – measures anhedonia, sleep disturbance, agitation, anxiety, weight loss
 - HRSD-RV – measures reverse vegetative symptoms such as weight gain and hyperphagia
- **Covariates:** treatment group, lifetime depression spectrum



The Model

$$y_{i1}(t_{ij}) = \mathbf{x}'_{ij}\boldsymbol{\mu}_1(t_{ij}) + \mathbf{z}'_{ij}\mathbf{g}_{i1}(t_{ij}) + \delta_{i1}(t_{ij})$$

$$y_{i2}(t_{ij}) = \mathbf{x}'_{ij}\boldsymbol{\mu}_2(t_{ij}) + \mathbf{z}'_{ij}\mathbf{g}_{i2}(t_{ij}) + \delta_{i2}(t_{ij})$$

$$\vdots = \vdots$$

$$y_{ip}(t_{ij}) = \mathbf{x}'_{ij}\boldsymbol{\mu}_p(t_{ij}) + \mathbf{z}'_{ij}\mathbf{g}_{ip}(t_{ij}) + \delta_{ip}(t_{ij})$$

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- $\boldsymbol{\mu}_k(t) = (\mu_{k1}(t), \dots, \mu_{kr}(t))'$ is an r -vector of fixed functions,
 $k = 1, \dots, p$

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- $\boldsymbol{\mu}_k(t) = (\mu_{k1}(t), \dots, \mu_{kr}(t))'$ is an r -vector of **fixed functions**, $k = 1, \dots, p$
- $\mathbf{g}_{ik}(t) = (g_{ik1}(t), \dots, g_{iks}(t))'$ is an s -vector of **random functions**, $k = 1, \dots, p$

The Model (Cont.)

$$\mathbf{y}_i(t_{ij}) = X_{ij}\boldsymbol{\mu}(t_{ij}) + Z_{ij}\mathbf{g}_i(t_{ij}) + \boldsymbol{\delta}_i(t_{ij}) ,$$

- $\mathbf{y}_i(t_{ij}) = (y_{i1}(t_{ij}), \dots, y_{ip}(t_{ij}))'$, $i = 1, \dots, n$, $j = 1, \dots, m_i$.

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- $X_{ij} = I_p \otimes \mathbf{x}'_{ij}$, $Z_{ij} = I_p \otimes \mathbf{z}'_{ij}$.
- $\boldsymbol{\delta}_i(t_{ij}) = (\delta_{i1}(t_{ij}), \dots, \delta_{ip}(t_{ij}))'$

Example

Assume 2 groups of subjects and a bivariate response ($p = 2$):

$$\mathbf{x}_{ij} = (x_{ij1}, x_{ij2})', \text{ where } x_{ij1} = 1, \quad x_{ij2} = \begin{cases} 1 & \text{if treatment} \\ 0 & \text{if control} \end{cases}$$

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$$y_{i1}(t_{ij}) = (1 \quad x_{ij2}) \begin{pmatrix} \mu_{11}(t_{ij}) \\ \mu_{12}(t_{ij}) \end{pmatrix} + g_{i1}(t_{ij}) + \delta_{i1}(t_{ij})$$

$$y_{i2}(t_{ij}) = (1 \quad x_{ij2}) \begin{pmatrix} \mu_{21}(t_{ij}) \\ \mu_{22}(t_{ij}) \end{pmatrix} + g_{i2}(t_{ij}) + \delta_{i2}(t_{ij})$$

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so

$$y_{i1}(t_{ij}) = \mu_{11}(t_{ij}) + x_{ij2} \cdot \mu_{12}(t_{ij}) + g_{i1}(t_{ij}) + \delta_{i1}(t_{ij})$$

$$y_{i2}(t_{ij}) = \mu_{21}(t_{ij}) + x_{ij2} \cdot \mu_{22}(t_{ij}) + g_{i2}(t_{ij}) + \delta_{i2}(t_{ij})$$

The Error Term

$\delta_i(t)$ follows a multivariate Ornstein-Uhlenbeck (O-U) process, i.e.,

$$d\delta_i(t) = -A\delta_i(t) + Bd\mathbf{W}_i(t) .$$

- A and B are $p \times p$ matrices of full rank common to all $i = 1, \dots, n$,
- $\mathbf{W}_i(t)$ is the p -dimensional Wiener process.
- Stationary if the EV of A have positive real parts.
- The stationary variance-covariance matrix Σ is the solution to

$$A\Sigma + \Sigma A' = BB'$$

The Error Term (Cont.)

In the **stationary** state, for $s < t$

$$\text{Cov}(\boldsymbol{\delta}_i(t), \boldsymbol{\delta}_i(s)) = \exp\{-A(t - s)\}\Sigma ,$$

so A controls how rapidly the influence of time s dies off by time $t > s$.

The Transition Density

Let $\Delta t_{ij} = t_{ij} - t_{i,j-1}$ $j = 1, \dots, m_i$, $t_{i0} = 0$.

The transition density of the O-U process is given by

$$p(\boldsymbol{\delta}_i(t_{ij}) | \boldsymbol{\delta}_i(t_{i,j-1}), \Delta t_{ij}) \propto |\Omega_{\Delta t_{ij}}|^{-1/2} \exp \left\{ -\frac{1}{2} \boldsymbol{\gamma}'_{t_{ij}} \Omega_{\Delta t_{ij}}^{-1} \boldsymbol{\gamma}_{t_{ij}} \right\},$$

where

$$\begin{aligned} \boldsymbol{\gamma}_{t_{ij}} &= \boldsymbol{\delta}_i(t_{ij}) - \exp(-A\Delta t_{ij})\boldsymbol{\delta}_i(t_{i,j-1}) \\ \Omega_{\Delta t_{ij}} &= \Sigma - \exp(-A\Delta t_{ij})\Sigma \exp(-A'\Delta t_{ij}). \end{aligned}$$

Nonparametric Function Estimation– Simple Example

200 observations were generated from

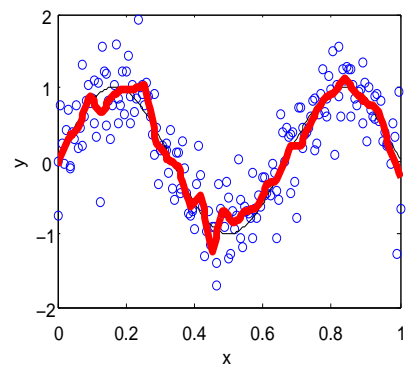
$$y_i = \sin(3\pi x_i) + \epsilon_i, \quad 0 \leq x_i \leq 1, \quad \epsilon_i \sim N(0, 0.4^2)$$

Model:

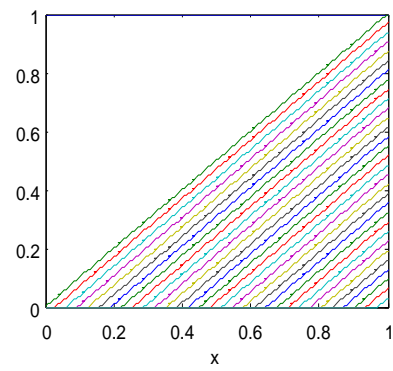
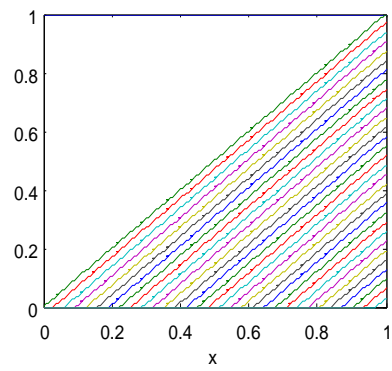
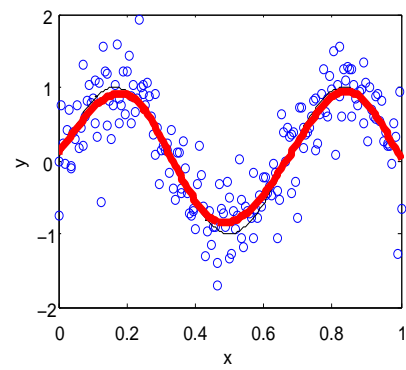
$$y_i = \beta_0 + \beta_1 x_i + \sum_{k=1}^{30} u_k (x_i - \kappa_k)_+ + \epsilon_i, \quad \epsilon_i \sim N(0, \sigma_\epsilon^2)$$

- Low rank truncated line basis functions.
- Fixed effects: β s and u s fixed.
- Mixed effects: β s fixed, $u_1, \dots, u_{30} \stackrel{iid}{\sim} N(0, \sigma_u^2)$.

Fixed Effects Model



Mixed Model



Nonparametric Function Estimation (Cont.)

We estimate

$\mu_{kl}(t)$ and $g_{ikm}(t)$, $k = 1, \dots, p$, $l = 1, \dots, r$, $m = 1, \dots, s$, $i = 1, \dots, n$

by **cubic splines** using **low-rank radial basis functions** (French et al., 2001, Ruppert et al., 2003):

Write $\mu_{kl}(t_{ij}) = \boldsymbol{\phi}'_{ij} \boldsymbol{\beta}_{kl} + \boldsymbol{\psi}'_{ij} \mathbf{v}_{kl}$ and $g_{ikm}(t_{ij}) = \boldsymbol{\phi}'_{ij} \mathbf{w}_{ikm} + \boldsymbol{\psi}'_{ij} \mathbf{u}_{ikm}$.

Nonparametric Function Estimation (Cont.)

We estimate

$$\mu_{kl}(t) \quad \text{and} \quad g_{ikm}(t), \quad k = 1, \dots, p, l = 1, \dots, r, m = 1, \dots, s, i = 1, \dots, n$$

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- Place K **knots** $\kappa_1, \dots, \kappa_K$ at sample quantiles of t_{ij} .

Nonparametric Function Estimation (Cont.)

We estimate

$$\mu_{kl}(t) \quad \text{and} \quad g_{ikm}(t), k = 1, \dots, p, l = 1, \dots, r, m = 1, \dots, s, i = 1, \dots, n$$

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- Place K **knots** $\kappa_1, \dots, \kappa_K$ at sample quantiles of t_{ij} .
- Let $\boldsymbol{\phi}'_{ij} = (1 \quad t_{ij})$, $\boldsymbol{\xi}'_{ij} = (|t_{ij} - \kappa_1|^3, \dots, |t_{ij} - \kappa_K|^3)$,
- Compute $\Lambda_K = [|\kappa_k - \kappa_{k'}|^3]_{1 \leq k, k' \leq K}$ and $\boldsymbol{\psi}'_{ij} = \boldsymbol{\xi}'_{ij} \Lambda_K^{-1/2}$ (via SVD)

Prior Distributions

Priors on the **basis function coefficients**:

1. Flat priors on β_{kl} .

2. $\mathbf{v}_{kl} \stackrel{ind}{\sim} N(\mathbf{0}, \sigma_{v_{kl}}^2 I_K)$

3. $\mathbf{w}_{ikm} \stackrel{ind}{\sim} N(\mathbf{0}, \text{diag}(\sigma_{w_{km0}}^2, \sigma_{w_{km1}}^2))$.

4. $\mathbf{u}_{ikm} \stackrel{ind}{\sim} N(\mathbf{0}, \sigma_{u_{km}}^2 I_K)$.

Priors on the **variance components**:

independent $IG(a_i, b_i)$, e.g., $p(\sigma_{v_{kl}}^2) \propto (\sigma_{v_{kl}}^2)^{-(a_1+1)} \exp(-b_1/\sigma_{v_{kl}}^2)$

Prior Distributions (Cont.)

Prior on A :

- Let $A = S\Psi S^{-1}$, where

$$S = \begin{pmatrix} 1 & s_{12} & s_{13} & \dots & s_{1p} \\ s_{21} & 1 & s_{23} & \dots & s_{2p} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ s_{p1} & s_{p2} & s_{p3} & \dots & 1 \end{pmatrix}$$

and $\Psi = \text{diag}(\psi_1, \dots, \psi_p)$, $\psi_i > 0$, $i = 1, \dots, p$.

- $s_{ij} \sim N(0, \sigma_s^2)$, $\log \psi_i \sim N(0, \sigma_\psi^2)$

Prior Distributions (Cont.)

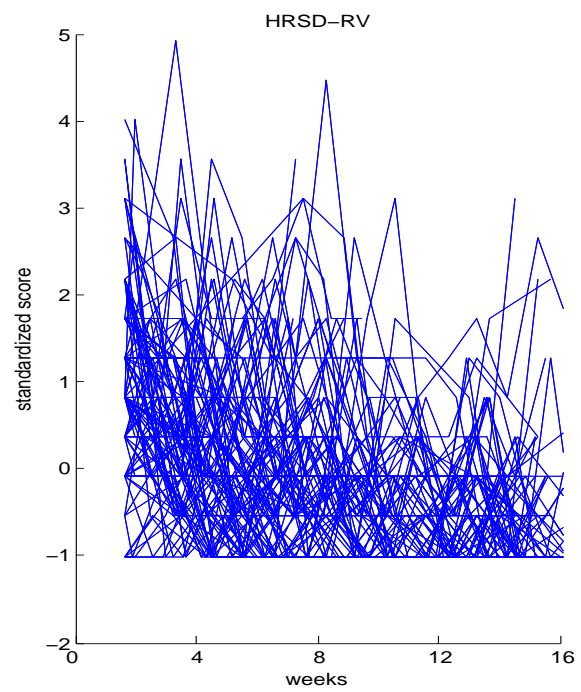
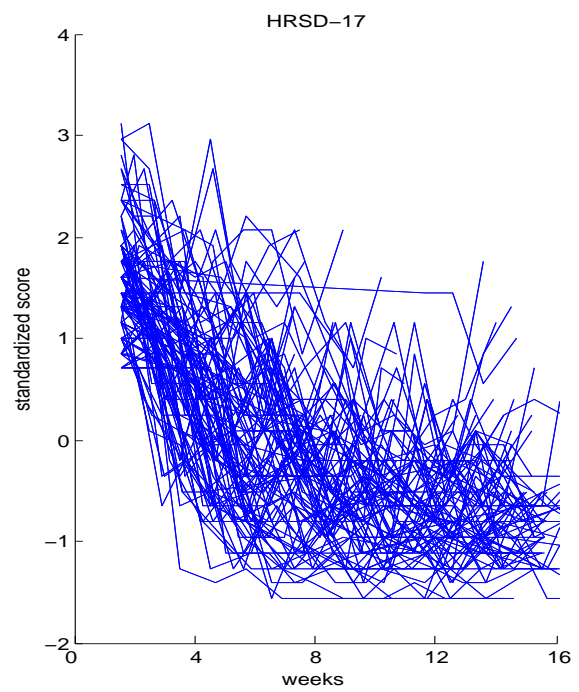
Prior on $C = BB'$:

- Let $C = LDL'$ (modified Cholesky factorization), where

$$L = \begin{pmatrix} 1 & 0 & 0 & \dots & 0 \\ l_{21} & 1 & 0 & \dots & 0 \\ l_{31} & l_{32} & 1 & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ l_{p1} & l_{p2} & l_{p3} & \dots & 1 \end{pmatrix}$$

and $D = \text{diag}(d_1, \dots, d_p)$, $d_i > 0$, $i = 1, \dots, p$

- $l_{ij} \sim N(0, \sigma_L^2)$, $\log(d_i) \sim N(0, \sigma_D^2)$



Example

- Psychiatric study to compare psychotherapy to pharmacotherapy at the University of Pisa, Italy. Total of 103 acutely depressed subjects.
- **Response variables:** 2 subscales of the clinician-administered Hamilton Depression rating Scale (measured at baseline and weekly):
 - HRSD-17 (0–31)
 - HRSD-RV (0–13).Higher values indicate more depressive symptoms.
- **Covariates:**
 - treatment group
 - lifetime depression spectrum (LDS), assessed at baseline.

Example (Cont.)

$$\mathbf{y}_i(t_{ij}) = X_{ij}\boldsymbol{\mu}(t_{ij}) + Z_{ij}\mathbf{g}_i(t_{ij}) + \boldsymbol{\delta}_i(t_{ij}) ,$$

where $X_{ij} = I_p \otimes \mathbf{x}'_{ij}$, $Z_{ij} = I_p \otimes \mathbf{z}'_{ij}$.

In the example $\mathbf{y}_i(t_{ij}) = (\text{HRSD-17}, \text{HRSD-RV})'$

$$\mathbf{z}_{ij} = 1$$

$$\mathbf{x}_{ij} = (x_{ij1}, x_{ij2}, x_{ij3}, x_{ij4})'$$

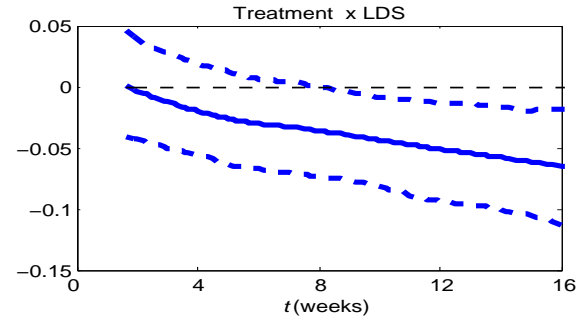
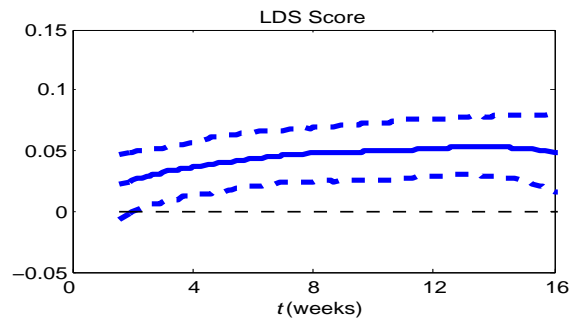
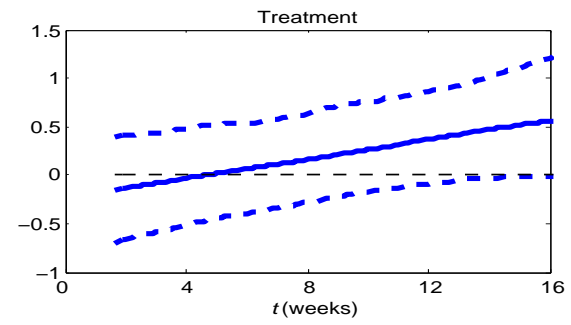
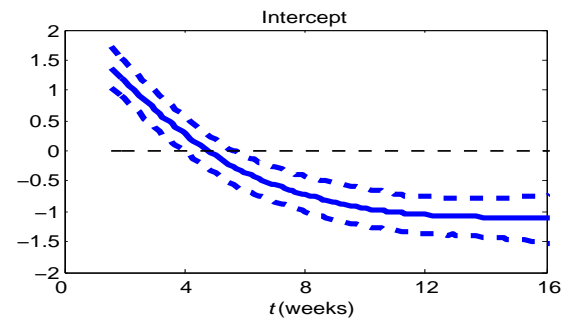
where $x_{ij1} = 1$

$$x_{ij2} = \begin{cases} 1 & \text{if subject } i \text{ received psychotherapy} \\ 0 & \text{if subject } i \text{ received pharmacotherapy,} \end{cases}$$

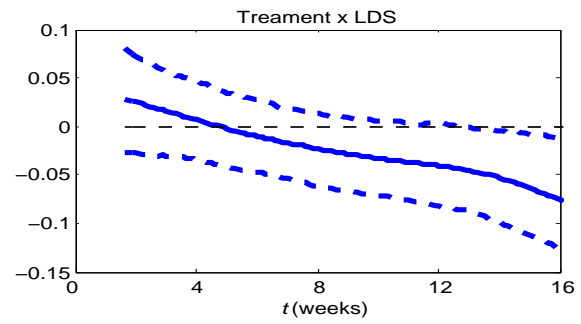
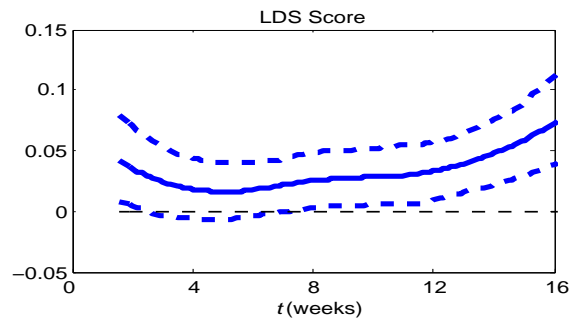
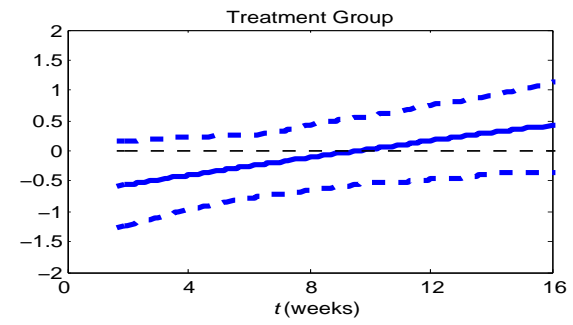
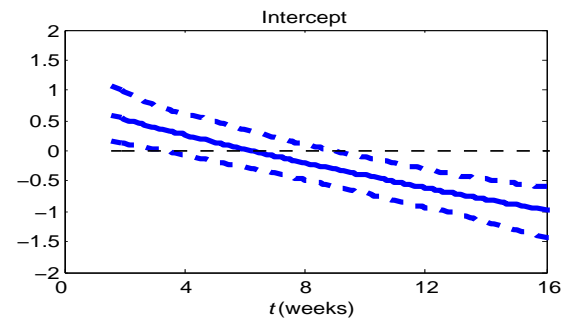
$$x_{ij3} = \text{LDS score}$$

$$x_{ij4} = x_{ij2}x_{ij3}.$$

HRSD-17



HRSD-RV



Conclusions from the Analysis

- Little evidence of a treatment effect
- Significant effect of LDS score – higher LDS predicts worse outcomes
- LDS effect is more pronounced in the middle period of HRSD-17, and towards the end of the 16 weeks for HRSD-RV.
- Interaction is significant in the last 8 weeks of HRSD-17, and marginally significant in the same time period for HRSD-RV.
- Interaction effect – LDS is less predictive of poor response in psychotherapy group than in pharmacotherapy group.
- Univariate fitting – interaction is not significant; wider pointwise credible intervals.