

Nonparametric Bayesian Kernel Models

Feng Liang

Joint work with Ming Liao, Sayan Mukherjee, Mike West

Institute of Statistics and Decision Sciences
Duke University

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Outline

1 Introduction

2 A Nonparametric Bayesian Approach

3 Semi-supervised learning

4 Conclusion

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- Assumptions: $(\mathbf{x}_i, y_i)_{i=1}^n \sim \mathcal{P}(\mathbf{x}, y)$.
- Find a function $f(\mathbf{x}) \rightarrow y$.
- Risk : $R[f] = \mathbb{E}L(Y, f(\mathbf{X}))$

$$f^*(\mathbf{x}) = \operatorname{argmin}_f R[f]$$

e.g., $f^*(\mathbf{x}) = \mathbb{E}(Y \mid \mathbf{X} = \mathbf{x})$ under squared error loss.

- Empirical risk minimization

$$\hat{f}_n(\mathbf{x}) = \operatorname{argmin}_{f \in \mathcal{F}} \frac{1}{n} \sum_{i=1}^n L(y_i, f(\mathbf{x}_i))$$

Regularization

- Regularization:

$$\min_{f \in \mathcal{F}} \hat{R}_n[f] + \lambda \Omega[f]$$

- Consider a linear regression model: $f(\mathbf{x}) = \mathbf{x}^t \boldsymbol{\beta}$,

$$\hat{\boldsymbol{\beta}}_{OLS} = \operatorname{argmin}_{\boldsymbol{\beta} \in \mathbb{R}^p} \sum_i (y_i - \mathbf{x}_i^t \boldsymbol{\beta})^2.$$

- Large p small n : $p \gg n$.
- $\Omega(\boldsymbol{\beta}) = \|\boldsymbol{\beta}\|^2$ (ridge); $|\boldsymbol{\beta}|$ (LASSO).

Reproducing Kernel Hilbert Space (RKHS)

- $\mathcal{H}_K = \{f(\mathbf{x}), \mathbf{x} \in \mathcal{X}\}$.
- $K(\cdot, \cdot)$ is a semi-positive definite bivariate symmetric function defined on $\mathcal{X} \times \mathcal{X}$, i.e.

$$\sum_{i=1}^m \sum_{j=1}^m a_i a_j K(\mathbf{x}_i, \mathbf{x}_j) \geq 0, \quad \forall a_i \in \mathbb{R}, \forall \mathbf{x}_i \in \mathcal{X}, \forall m \in \mathbb{N}.$$

- Denote $k_{\mathbf{u}}(\cdot) = K(\mathbf{u}, \cdot)$.

$$\mathcal{H}_K = \overline{\text{span}\{k_{\mathbf{u}}(\cdot), \mathbf{u} \in \mathcal{X}\}}.$$

- Reproducing property: $\forall f \in \mathcal{H}_K, \forall \mathbf{x}_0 \in \mathcal{X}$

$$f(\mathbf{x}_0) = \langle f, k_{\mathbf{x}_0} \rangle = \langle f, K(\mathbf{x}_0, \cdot) \rangle.$$

A Simple Examples of RKHS

- Consider all lineal functions in \mathbb{R}^2 passing the origin, i.e.

$$\mathcal{H}_K = \{f_\theta(\mathbf{x}) = \theta^t \mathbf{x} = \theta_1 x_1 + \theta_2 x_2, \theta \in \mathbb{R}^2\}$$

with

$$\langle f_\theta, f_\lambda \rangle = \theta^t \lambda, \quad K(\mathbf{x}, \mathbf{x}') = \mathbf{x}^t \mathbf{x}' = x_1 x'_1 + x_2 x'_2.$$

- Reproducing property : $\forall f_\theta,$

$$f_\theta(\mathbf{x}_0) = \theta^t \mathbf{x}_0 = \langle f_\theta, f_{\mathbf{x}_0} \rangle.$$

- Semi-positive definite :

$$\sum_i \sum_j a_i a_j K(\mathbf{x}_i, \mathbf{x}_j) = \left(\sum_i a_i \mathbf{x}_i \right)^t \cdot \left(\sum_j a_j \mathbf{x}_j \right) = \left\| \sum_i a_i \mathbf{x}_i \right\|^2$$

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- Any RKHS is a space of *linear* functions and $K(\cdot, \cdot)$ represents a dot product in the **feature space**.
- Feature mapping $\phi: \mathbf{x} \in \mathbb{R}^p \longrightarrow \phi(\mathbf{x}) \in \mathbb{R}^N$.

$$(x_1, x_2) \in \mathbb{R}^2 \longrightarrow (1, x_1, x_2, x_1^2, x_2^2, x_1 x_2) \in \mathbb{R}^6$$

- $\mathcal{H}_K = \{f_\theta(\mathbf{x}) = \theta^t \cdot \phi(\mathbf{x})\}$ and $K(\mathbf{x}, \mathbf{x}') = \phi(\mathbf{x})^t \cdot \phi(\mathbf{x}')$.
- Examples
 - Polynomial kernels: $K(\mathbf{x}, \mathbf{x}') = (1 + \mathbf{x}^t \mathbf{x}')^d$.
 - Gaussian radial basis kernels:

$$K(\mathbf{x}, \mathbf{x}') = \exp \left(- \frac{\|\mathbf{x} - \mathbf{x}'\|^2}{2\sigma^2} \right).$$

Representer Theorem

(Kimeldorf and Wahba, 1971)

$$f^*(\mathbf{x}) = \operatorname{argmin}_{f \in \mathcal{H}_K} \frac{1}{n} \sum_{i=1}^n L(y_i, f(\mathbf{x}_i)) + \lambda \|f\|_{\mathcal{H}_K}^2$$

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Proof : Let $\mathcal{H}_1 = \text{span}\{k_{\mathbf{x}_1}, \dots, k_{\mathbf{x}_n}\}$ and $\mathcal{H}_2 = \mathcal{H}_1^\perp$,

$$f = f_1 + f_2, \quad \mathcal{H}_K = \mathcal{H}_1 \oplus \mathcal{H}_2$$

- $\|f\|^2 \geq \|f_1\|^2$
- $f(\mathbf{x}_i) = f_1(\mathbf{x}_i)$ because

$$\langle f, k_{\mathbf{x}_i} \rangle = \langle f_1 + f_2, k_{\mathbf{x}_i} \rangle = \langle f_1, k_{\mathbf{x}_i} \rangle$$

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Primer on Bayesian Analysis

- Data and a parametric family: $\mathbf{y} \sim p(\cdot \mid \theta)$

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Primer on Bayesian Analysis

- Data and a parametric family: $\mathbf{y} \sim p(\cdot | \theta)$
- Prior on θ : $\pi(\theta)$
- Joint distribution on (\mathbf{y}, θ) : $p(\mathbf{y} | \theta)\pi(\theta)$
- Posterior inference:

$$\pi(\theta | \mathbf{y}) = \frac{p(\mathbf{y} | \theta)\pi(\theta)}{\int p(\mathbf{y} | \theta)\pi(\theta)d\theta} \propto p(\mathbf{y} | \theta)\pi(\theta)$$

Connection to Regularization

- Log posterior = $\log p(\mathbf{y} \mid \theta) + \log \pi(\theta) + \dots$

$$\text{Posterior Mode} = \operatorname{argmin}_{\theta} \sum_i (y_i - f_{\theta}(\mathbf{x}_i))^2 - c \log \pi(\theta)$$

- Regularization

$$\operatorname{argmin}_{\theta} \sum_{i=1} (y_i - f_{\theta}(\mathbf{x}_i))^2 + \lambda \Omega[f]$$

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$$\text{Posterior Mode} = \operatorname{argmin}_i (y_i - f_\theta(\mathbf{x}_i))^2 - c \log \pi(\theta)$$

- Regularization

$$\operatorname{argmin}_{i=1} (y_i - f(\mathbf{x}_i))^2 + \lambda \Omega[f]$$

- $\pi(\theta) \propto \exp\{-\lambda \Omega[f]\}$. For example, when $f_\theta = \theta^t \mathbf{x}$,
 - Ridge ($\|\theta\|^2$) \iff normal
 - LASSO ($|\theta|$) \iff double exponential

Previous Work

For example, Tipping 2001, Chakraborty et al. 2005, and others.

- Start with the **finite** representation from the representer Theorem:

$$\sum_{i=1}^n \alpha_i K(\mathbf{x}, \mathbf{x}_i) \quad (*)$$

- Specify priors on the coefficients α_i 's.
- Their models change when sample size changes **without** a coherent argument.
- Can we justify $(*)$ using the connection between regularization and posterior mode?

An Orthonormal Representation of \mathcal{H}_K

- For Mercer kernels,

$$K(\mathbf{x}, \mathbf{u}) = \sum_{j=1}^{\infty} \lambda_j \phi_j(\mathbf{x}) \phi_j(\mathbf{u}),$$

where ϕ_j is a sequence of orthonormal functions and $\lambda_1 \geq \lambda_2 \geq \dots \geq 0$.

- (Cucker and Smale, 2001) $\forall f \in \mathcal{H}_K$,

$$f(\mathbf{x}) = \sum_j \alpha_j \phi_j(\mathbf{x}), \quad \sum_j \frac{\alpha_j^2}{\lambda_j} < \infty.$$

That is, \mathcal{H}_K can be parameterized by

$$\mathcal{A} = \{(\alpha_j)_{j=1}^{\infty} : \sum_j \alpha_j^2 / \lambda_j < \infty\}.$$

An Overcomplete Representation of \mathcal{H}_K

- Recall

$$\mathcal{H}_K = \overline{\text{span}\{K(\cdot, \mathbf{u}), \mathbf{u} \in \mathcal{X}\}}$$

- Start with a larger space

$$f(\mathbf{x}) = \int_{\mathcal{X}} K(\mathbf{x}, \mathbf{u}) d\Gamma(\mathbf{u}),$$

where $\Gamma(\mathbf{u})$ is a sign measure on \mathcal{X} .

- In this talk, we focus on the following representation

$$f(\mathbf{x}) = \int_{\mathcal{X}} w(\mathbf{u}) K(\mathbf{x}, \mathbf{u}) dF(\mathbf{u}),$$

where $w(\mathbf{u})$ denotes the coefficient at location \mathbf{u} and F denotes the distribution function of the location \mathbf{u} .

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$$f(\mathbf{x}) = \int_{\mathcal{X}} w(\mathbf{u}) K(\mathbf{x}, \mathbf{u}) dF_{\mathbf{X}}(\mathbf{u}),$$

where $w(\mathbf{u})$ denotes the coefficient at location \mathbf{u} and $F_{\mathbf{X}}$ denotes the distribution function of explanatory variable \mathbf{X} .

Dirichlet Process Priors

- Beta (α, β) on $x \in [0, 1]$

$$f(x) \propto x^{\alpha-1} (1-x)^{\beta-1}$$

- Dir $(\alpha_1, \dots, \alpha_k)$ on (x_1, \dots, x_k) where $x_i \in [0, 1]$ and $\sum_i x_i = 1$

$$f(x) \propto x_1^{\alpha_1-1} \cdots x_k^{\alpha_k-1}$$

- DP(α_0, F_0) on \mathbf{F} (note \mathbf{F} is a random distribution on \mathcal{X}):
for any measurable partition of \mathcal{X} ,

$$\mathbf{F}(B_1), \mathbf{F}(B_2), \dots, \mathbf{F}(B_k) \sim \text{Dir}(\alpha_1, \dots, \alpha_k),$$

where $\alpha_i = \alpha_0 F_0(B_i)$.

A Bayesian Representer Theorem

Given $\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n \sim \mathbf{F}$, with $\text{DP}(\alpha_0, F_0)$ prior on \mathbf{F} , the posterior distribution of \mathbf{F} is $\text{DP}(\alpha_0 + n, \alpha_0 F_0 + \sum \delta_{\mathbf{x}_i})$. So

$$\mathbb{E}[\mathbf{F} \mid \mathbf{x}_1, \dots, \mathbf{x}_n] = \frac{\alpha_0}{\alpha_0 + n} F_0 + \frac{1}{\alpha_0 + n} \sum_{i=1}^n \delta_{\mathbf{x}_i}.$$

Bayesian Representer Theorem

For $f(\mathbf{x}) = \int w(\mathbf{u}) K(\mathbf{x}, \mathbf{u}) d\mathbf{F}_X(\mathbf{u})$, under a Dirichlet prior on \mathbf{F}_X ,

$$\mathbb{E}[f(\mathbf{x})] \approx \sum_{i=1}^n w_i K(\mathbf{x}, \mathbf{x}_i).$$

Prior Specification

- Likelihood: $\mathbf{Y} \mid \mathcal{N}(\mathbf{w}_0 + \mathbf{K}\mathbf{w}, \sigma^2 I_n)$ where $\mathbf{K}_{n \times n}$ be the centered kernel matrix at the n data points.
- Non-informative prior on (\mathbf{w}_0, σ^2) ,

$$\pi(w_0, \sigma^2) \propto 1/\sigma^2$$

- Generalized g-prior on \mathbf{w} (West 2002)

$$\begin{aligned} \mathbf{w} \mid \mathbf{T} &\sim \mathcal{N}(0, U\Delta^{-1}T^{-1}\Delta^{-1}U^t) \\ \tau_1, \dots, \tau_n &\sim \text{Gamma}\left(\frac{s_0}{2}, \frac{s_0}{2}v\right), \quad v \sim \text{Exp}(\alpha_0) \end{aligned}$$

where U and Δ come from $\mathbf{K} = U\Delta U^t$.

- $\mathbf{K}\mathbf{w} = U\beta$, then the prior on \mathbf{w} corresponds to a student t distribution on β .

Model Fitting via MCMC

- Gibbs sampling
 - $w_0 = \dots$
 - Draw $\beta \sim \mathcal{N}(\cdot, \cdot)$
 - Draw $T = \text{diag}(\tau_1, \dots, \tau_n) \sim Ga(\cdot, \cdot)$
 - Draw $v \sim Ga(\cdot, \cdot)$
- Probit model: $(\mathbf{x}, y) \longrightarrow (\mathbf{x}, y, \mathbf{z})$ and

$$P(y = 1) = \Phi(z).$$

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Semi-supervised Learning

- Supervised learning (**labelled data**):

$$(\mathbf{x}_i, y_i) \sim P(\mathbf{x}, y)$$

- Unsupervised learning (**unlabelled data**):

$$(\mathbf{x}_i) \sim P_{\mathbf{X}}(\mathbf{x})$$

- Semi-supervised learning:

$$(\mathbf{x}_i, y_i)_{i=1}^n \sim P, \quad (\mathbf{x}_i)_{i=n+1}^{n+m} \sim P_{\mathbf{X}}$$

- How's it different from missing data?

The Role of Unlabelled Data

- Data $D = (\mathbf{X}, Y, \mathbf{X}^U)$ where

$$\begin{aligned}(\mathbf{X}, Y) &\sim P(\mathbf{x}, y) = p(\mathbf{x} \mid \phi) p(y \mid \mathbf{x}, \theta) \\ \mathbf{X}^U &\sim P_{\mathbf{X}}(\mathbf{x}) = p(\mathbf{x} \mid \phi)\end{aligned}$$

- Prediction of y^* at a new location \mathbf{x}^* ,

$$\begin{aligned}\mathbb{E}[Y^* \mid x^*, D] &= \int y^* p(y^* \mid x^*, D) dy^* \\ &= \int y^* p(y^* \mid \mathbf{x}^*, \theta) \pi(\theta \mid D) d\theta dy^*.\end{aligned}$$

- The key to understand the role of \mathbf{X}^m is

$$\begin{aligned}\pi(\theta \mid D) &= \int \pi(\theta, \phi \mid D) d\phi \\ &\propto \int p(\mathbf{X}, \mathbf{X}^U \mid \phi) p(Y \mid \mathbf{X}, \theta) \pi(\theta, \phi) d\phi\end{aligned}$$

An Intimate Relationship

- $P(\mathbf{x}, y) = p(\mathbf{x} \mid \phi)p(y \mid \mathbf{x}, \theta)$
- Recall that $y \mid \mathbf{x} \sim \mathcal{N}(f(\mathbf{x}), \sigma^2)$ where

$$f(\mathbf{x}) = \int w(\mathbf{u})K(\mathbf{x}, \mathbf{u})dF_{\mathbf{X}}(\mathbf{u}).$$

So in our model, $\theta = (\phi, \dots)$. Therefore unlabelled data will be relevant and should be incorporated into prediction.

- By our Bayesian representor Theorem, given $(\mathbf{x}_i, y_i)_{i=1}^n$ and $(\mathbf{x}_j)_{j=n+1}^{n+m}$,

$$f(\mathbf{x}) \approx \sum_{i=1}^n w_i K(\mathbf{x}, \mathbf{x}_i) + \sum_{j=n+1}^{n+m} w_j K(\mathbf{x}, \mathbf{x}_j).$$

Connection to Regularization

- Transductive SVM (TSVM) (Joachims, 1999)

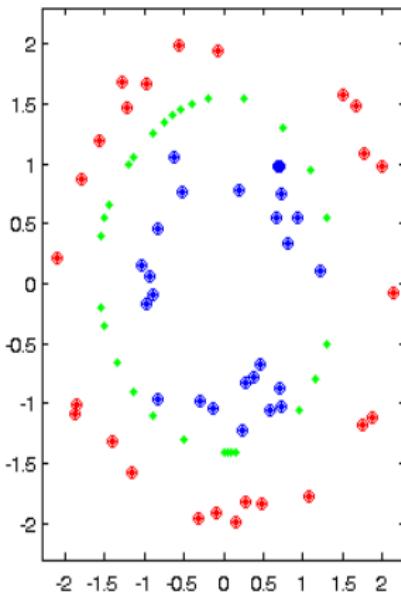
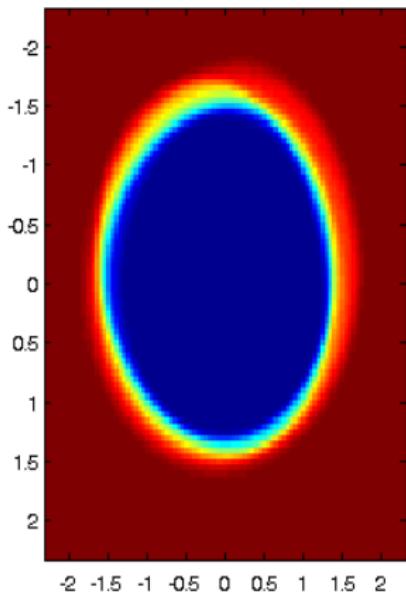
$$\operatorname{argmin} C \sum_{i=1}^n L(y_i, f(\mathbf{x}_i)) + C^* \sum_{i=n+1}^{n+m} L(y_i, f(\mathbf{x}_i)) + \|f\|_{\mathcal{H}_K}^2$$

- Manifold regularization (Belkin et al, 2005)

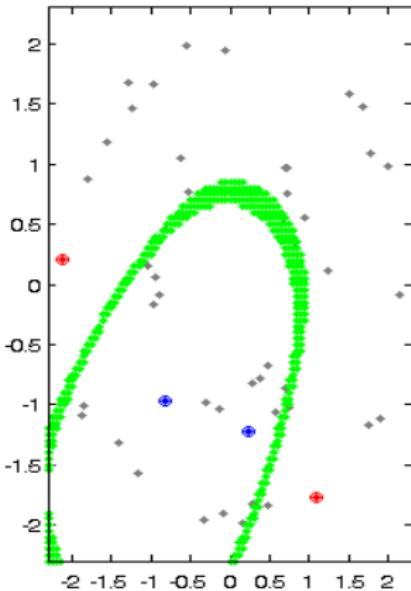
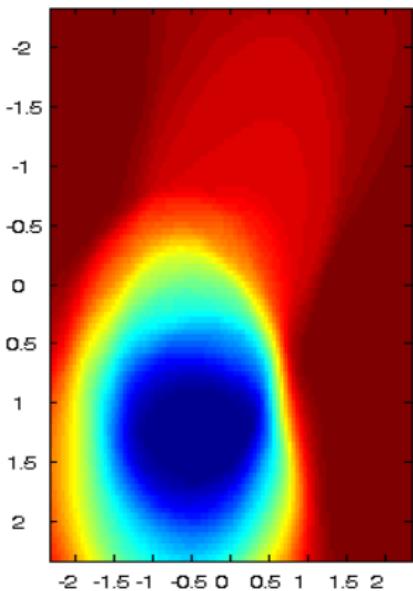
$$\operatorname{argmin} \sum_{i=1}^n L(y_i, f(\mathbf{x}_i)) + \lambda_1 \|f\|_{\mathcal{H}_K}^2 + \lambda_2 \|f\|_I^2,$$

where $\|f\|_I^2$ measures the intrinsic structure of F_X and is approximated on all the data (include the unlabelled ones) using graph Laplacian.

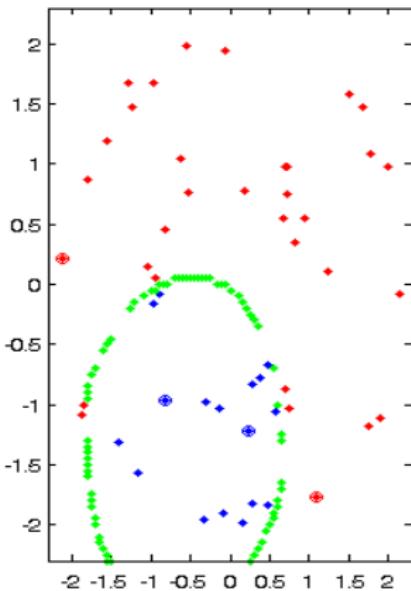
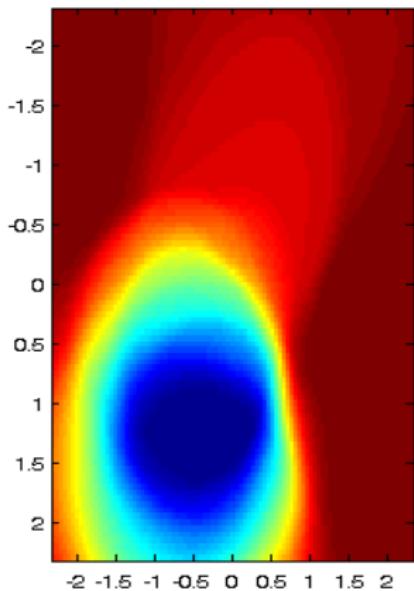
A Toydata



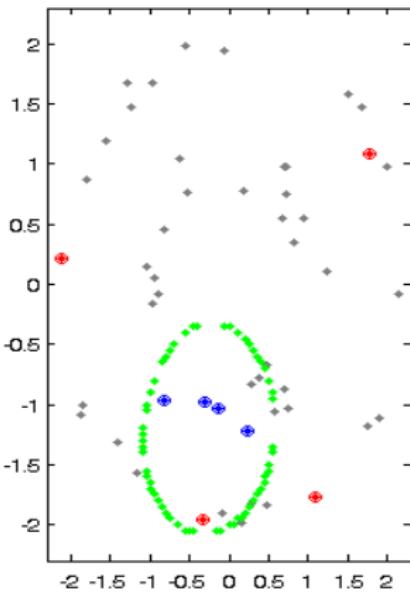
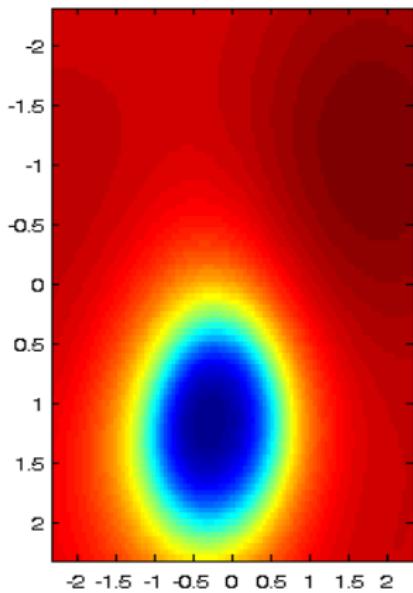
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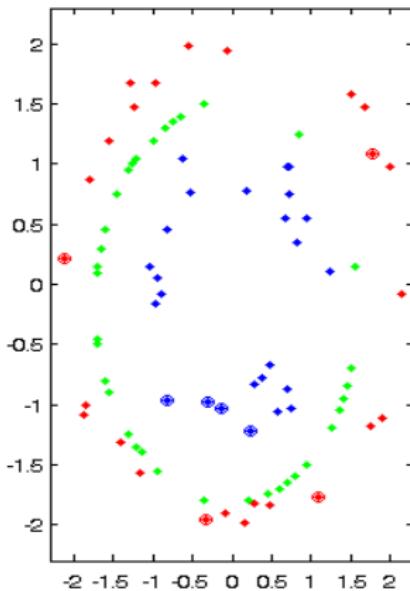
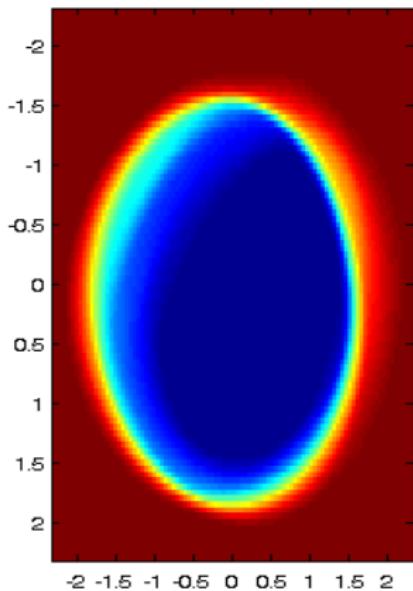
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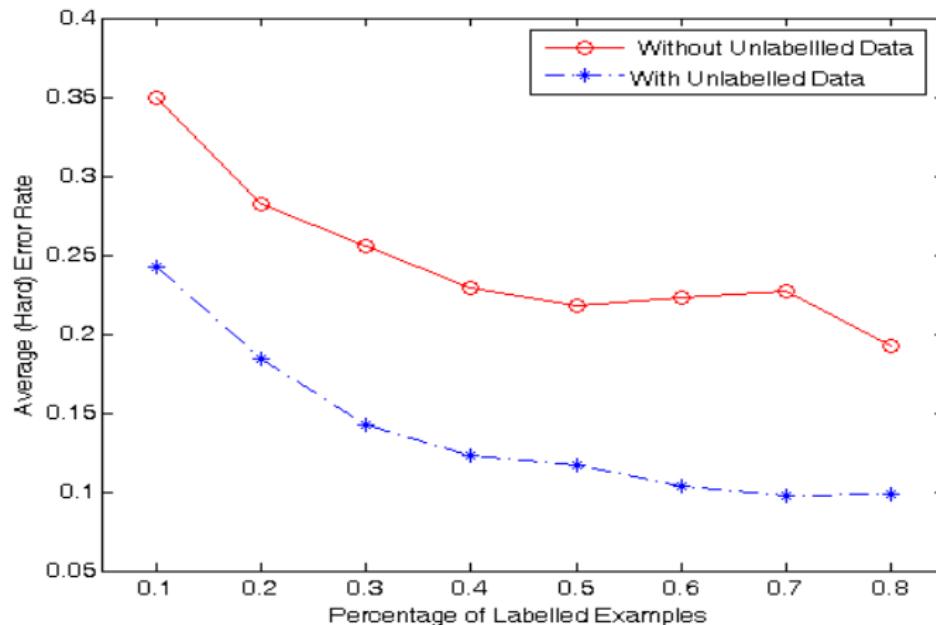
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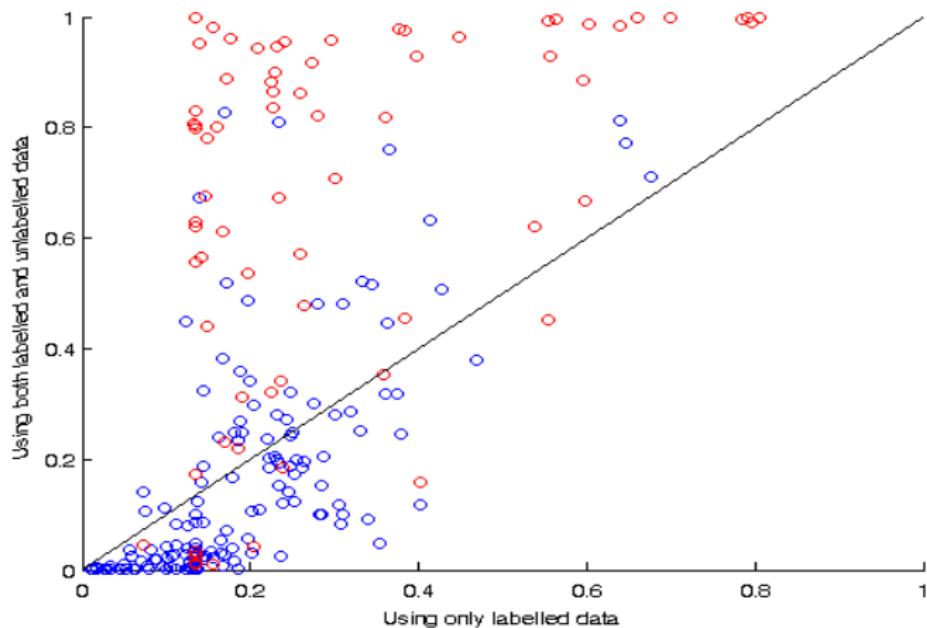
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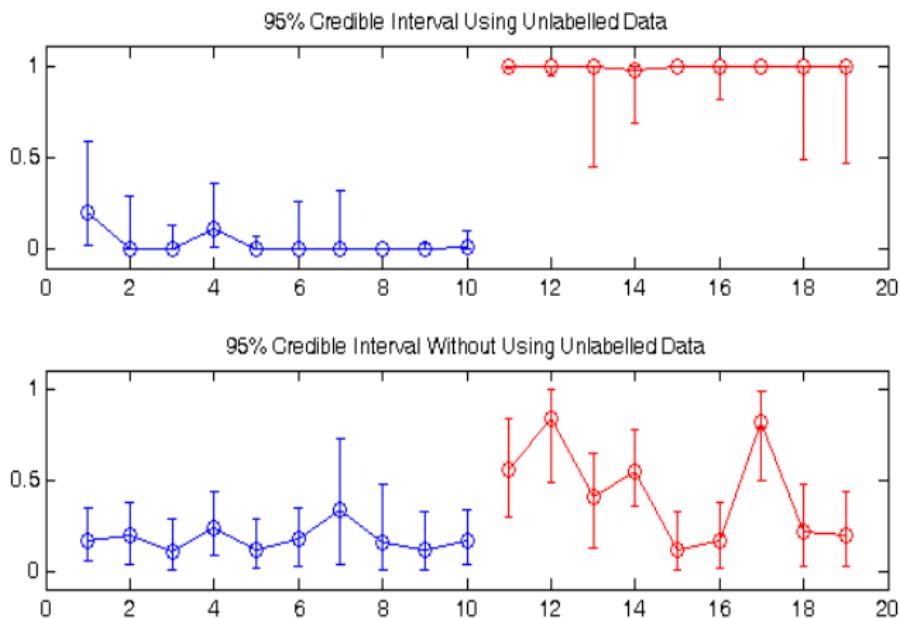
Cancer Data



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 - specify priors on the whole RKHS;
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- A coherent Bayesian approach on RKHS:
 - characterize RKHS using an overcomplete representation;
 - specify priors on the whole RKHS;
 - incorporate the relevant information from the unlabelled data.
- Future work:
 - Other choice of priors and sensitivity study
 - Feature selection