

Fractional hot deck imputation for multivariate missing data in survey sampling

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Features

- Split the record with missing item into $m(> 1)$ imputed values
- Assign fractional weights
- The final product is a single data file with size $\leq nm$.
- For variance estimation, the fractional weights are replicated.

Fractional imputation

Example ($n = 10$)

ID	Weight	y_1	y_2
1	w_1	$y_{1,1}$	$y_{1,2}$
2	w_2	$y_{2,1}$	M
3	w_3	M	$y_{3,2}$
4	w_4	$y_{4,1}$	$y_{4,2}$
5	w_5	$y_{5,1}$	$y_{5,2}$
6	w_6	$y_{6,1}$	$y_{6,2}$
7	w_7	M	$y_{7,2}$
8	w_8	M	M
9	w_9	$y_{9,1}$	$y_{9,2}$
10	w_{10}	$y_{10,1}$	$y_{10,2}$

M: Missing

Fractional Imputation Idea

If both y_1 and y_2 are categorical, then fractional imputation is easy to apply.

- We have only finite number of possible values.
- Imputed values = possible values
- The fractional weights are the conditional probabilities of the possible values given the observations.
- Can use “EM by weighting” method of Ibrahim (1990) to compute the fractional weights.

Fractional imputation

Example (y_1, y_2 : dichotomous, taking 0 or 1)

ID	Weight	y_1	y_2
1	w_1	$y_{1,1}$	$y_{1,2}$
2	$w_2 w_{2,1}^*$	$y_{2,1}$	0
	$w_2 w_{2,2}^*$	$y_{2,1}$	1
3	$w_3 w_{3,1}^*$	0	$y_{3,2}$
	$w_3 w_{3,2}^*$	1	$y_{3,2}$
4	w_4	$y_{4,1}$	$y_{4,2}$
5	w_5	$y_{5,1}$	$y_{5,2}$

Fractional imputation

Example (y_1, y_2 : dichotomous, taking 0 or 1)

ID	Weight	y_1	y_2
6	w_6	$y_{6,1}$	$y_{6,2}$
7	$w_7 w_{7,1}^*$	0	$y_{7,2}$
	$w_7 w_{7,2}^*$	1	$y_{7,2}$
8	$w_8 w_{8,1}^*$	0	0
	$w_8 w_{8,2}^*$	0	1
	$w_8 w_{8,3}^*$	1	0
	$w_8 w_{8,4}^*$	1	1
9	w_9	$y_{9,1}$	$y_{9,2}$
10	w_{10}	$y_{10,1}$	$y_{10,2}$

Fractional imputation

Example (Cont'd)

- E-step: Fractional weights are the conditional probabilities of the imputed values given the observations.

$$\begin{aligned}w_{ij}^* &= \hat{P}(y_{i,mis}^{*(j)} \mid y_{i,obs}) \\ &= \frac{\hat{\pi}(y_{i,obs}, y_{i,mis}^{*(j)})}{\sum_{l=1}^{M_i} \hat{\pi}(y_{i,obs}, y_{i,mis}^{*(l)})}\end{aligned}$$

where $(y_{i,obs}, y_{i,mis})$ is the (observed, missing) part of $y_i = (y_{i1}, \dots, y_{ip})$.

- M-step: Update the joint probability using the fractional weights.

$$\hat{\pi}_{ab} = \frac{1}{\hat{N}} \sum_{i=1}^n \sum_{j=1}^{M_i} w_i w_{ij}^* I(y_{i,1}^{*(j)} = a, y_{i,2}^{*(j)} = b)$$

with $\hat{N} = \sum_{i=1}^n w_i$.

Example (Cont'd) Variance estimation

- Recompute the fractional weights for each replication
- Apply the same EM algorithm using the replicated weights.
 - E-step: Fractional weights are the conditional probabilities of the imputed values given the observations.

$$w_{ij}^{*(k)} = \frac{\hat{\pi}^{(k)}(y_{i,obs}, y_{i,mis}^{*(j)})}{\sum_{l=1}^{M_i} \hat{\pi}^{(k)}(y_{i,obs}, y_{i,mis}^{*(l)})}$$

- M-step: Update the joint probability using the fractional weights.

$$\hat{\pi}_{ab}^{(k)} = \frac{1}{\hat{N}^{(k)}} \sum_{i=1}^n \sum_{j=1}^{M_i} w_i^{(k)} w_{ij}^{*(k)} I(y_{i,1}^{*(j)} = a, y_{i,2}^{*(j)} = b)$$

where $\hat{N}^{(k)} = \sum_{i=1}^n w_i^{(k)}$.

Fractional imputation

Example (Cont'd) Final Product

Weight	x	y ₁	y ₂	Replication Weights			
				Rep 1	Rep 2	...	Rep L
w ₁	x ₁	y _{1,1}	y _{1,2}	w ₁ ⁽¹⁾	w ₁ ⁽²⁾	...	w ₁ ^(L)
w ₂ w _{2,1} [*]	x ₂	y _{2,2}	0	w ₂ ⁽¹⁾ w _{2,1} ^{*(1)}	w ₂ ⁽²⁾ w _{2,1} ^{*(2)}	...	w ₂ ^(L) w _{2,1} ^{*(L)}
w ₂ w _{2,2} [*]	x ₂	y _{2,2}	1	w ₂ ⁽¹⁾ w _{2,2} ^{*(1)}	w ₂ ⁽²⁾ w _{2,2} ^{*(2)}	...	w ₂ ^(L) w _{2,2} ^{*(L)}
w ₃ w _{3,1} [*]	x ₃	0	y _{3,2}	w ₃ ⁽¹⁾ w _{3,1} ^{*(1)}	w ₃ ⁽²⁾ w _{3,1} ^{*(2)}	...	w ₃ ^(L) w _{3,1} ^{*(L)}
w ₃ w _{3,2} [*]	x ₃	1	y _{3,2}	w ₃ ⁽¹⁾ w _{3,2} ^{*(1)}	w ₃ ⁽²⁾ w _{3,2} ^{*(2)}	...	w ₃ ^(L) w _{3,2} ^{*(L)}
w ₄	x ₄	y _{4,1}	y _{4,2}	w ₄ ⁽¹⁾	w ₄ ⁽²⁾	...	w ₄ ^(L)
w ₅	x ₅	y _{5,1}	y _{5,2}	w ₅ ⁽¹⁾	w ₅ ⁽²⁾	...	w ₅ ^(L)
w ₆	x ₆	y _{6,1}	y _{6,2}	w ₆ ⁽¹⁾	w ₆ ⁽²⁾	...	w ₆ ^(L)

Fractional imputation

Example (Cont'd) Final Product

Weight	x	y ₁	y ₂	Replication Weights			
				Rep 1	Rep 2	Rep L	
$w_7 w_{7,1}^*$	x ₇	0	y _{7,2}	$w_7^{(1)} w_{7,1}^{*(1)}$	$w_7^{(2)} w_{7,1}^{*(2)}$	$w_7^{(L)} w_{7,1}^{*(L)}$	
$w_7 w_{7,2}^*$	x ₇	1	y _{7,2}	$w_7^{(1)} w_{7,2}^{*(1)}$	$w_7^{(2)} w_{7,2}^{*(2)}$	$w_7^{(L)} w_{7,2}^{*(L)}$	
$w_8 w_{8,1}^*$	x ₈	0	0	$w_8^{(1)} w_{8,1}^{*(1)}$	$w_8^{(1)} w_{8,1}^{*(2)}$	$w_8^{(L)} w_{8,1}^{*(L)}$	
$w_8 w_{8,2}^*$	x ₈	0	1	$w_8^{(1)} w_{8,2}^{*(1)}$	$w_8^{(2)} w_{8,2}^{*(2)}$	$w_8^{(L)} w_{8,2}^{*(L)}$	
$w_8 w_{8,3}^*$	x ₈	1	0	$w_8^{(1)} w_{8,3}^{*(1)}$	$w_8^{(2)} w_{8,3}^{*(2)}$	$w_8^{(L)} w_{8,3}^{*(L)}$	
$w_8 w_{8,4}^*$	x ₈	1	1	$w_8^{(1)} w_{8,4}^{*(1)}$	$w_8^{(1)} w_{8,4}^{*(2)}$	$w_8^{(L)} w_{8,4}^{*(L)}$	
w ₉	x ₉	y _{9,1}	y _{9,2}	$w_9^{(1)}$	$w_9^{(2)}$...	$w_9^{(L)}$
w ₁₀	x ₁₀	y _{10,1}	y _{10,2}	$w_{10}^{(1)}$	$w_{10}^{(2)}$...	$w_{10}^{(L)}$

Fractional hot deck imputation

Goals

- Fractional hot deck imputation of size m . The final product is a single data file with size $\leq n \cdot m$.
- Preserves correlation structure
- Variance estimation relatively easy
- Can handle domain estimation (but we do not know which domains will be used.)

Parametric model approach

- In the parametric model approach, imputed values can be generated from $f(y_{i,mis} | y_{i,obs}, \hat{\theta})$ where $\hat{\theta}$ is the MLE of θ .
- Kim (2011) proposed parametric fractional imputation (PFI) which is based on importance sampling idea

$$w_{ij}^* \propto \frac{f(y_{i,obs}, y_{i,mis}^{*(j)}; \hat{\theta})}{h(y_{i,mis}^{*(j)} | y_{i,obs})}$$

with $\sum_{j=1}^m w_{ij}^* = 1$, where $y_{i,mis}^{*(1)}, \dots, y_{i,mis}^{*(m)} \sim h(y_{i,mis} | y_{i,obs})$.

Fractional hot deck imputation

Idea

- In hot deck imputation, we can make a nonparametric approximation of $f(\cdot)$ using a finite mixture model

$$f(y_{i,mis} | y_{i,obs}) = \sum_{g=1}^G \pi_g(y_{i,obs}) f_g(y_{i,mis}), \quad (1)$$

where $\pi_g(y_{i,obs}) = P(z_i = g | y_{i,obs})$, $f_g(y_{i,mis}) = f(y_{i,mis} | z = g)$ and z is the latent variable associated with imputation cell.

- To satisfy the above approximation, we need to find z such that

$$f(y_{i,mis} | z_i, y_{i,obs}) = f(y_{i,mis} | z_i).$$

Fractional hot deck imputation

Imputation cell

- Assume p -dimensional survey items: $Y = (Y_1, \dots, Y_p)$
- For each item k , create a transformation of Y_k into Z_k , a discrete version of Y_k based on the sample quantiles among respondents.
- If $y_{i,k}$ is missing, then $z_{i,k}$ is also missing.
- Imputation cells are created based on the observed value of $z_i = (z_{i,1}, \dots, z_{i,p})$.
- Expression (1) can be written as

$$f(y_{i,mis} \mid y_{i,obs}) = \sum_{Z_{mis}} P(z_{i,mis} = Z_{mis} \mid z_{i,obs}) f(y_{i,mis} \mid Z_{mis}), \quad (2)$$

where $z_i = (z_{i,obs}, z_{i,mis})$ similarly to $y_i = (y_{i,obs}, y_{i,mis})$.

Estimation of cell probability

- Let $\{z_1, \dots, z_G\}$ be the support z , which is the same as the sample support of z from the full respondents.
- Cell probability $\pi_g = P(z = z_g)$.
- For each unit i , we only observe $z_{i,obs}$.
- Use EM algorithm for categorical missing data to estimate π_g .

Two-stage sampling for fractional hot deck imputation of size m

- Stage 1: Given $z_{i,obs}$, imputed m values of $z_{i,mis}$, denoted by $z_{i,mis}^{*(1)}, \dots, z_{i,mis}^{*(m)}$, from the estimated conditional probability $P(z_{i,mis} \mid z_{i,obs})$.
- Stage 2: For each imputed cell $z_i^{*(j)} = (z_{i,obs}, z_{i,mis}^{*(j)})$, the imputed value for $y_{i,mis}$ is randomly chosen among the full respondents in the same cell. (Joint hot deck within imputation cell)

Variance estimation

- Kim and Fuller (2004) proposed a variance estimation method for fractional imputation under the cell mean model.
- Kim, Fuller, and Bell (2011) applied the method to variance estimation for income estimates in the 2000 Census long form data.

Example : Jackknife for Fractional Imputation in Kim and Fuller (2004)

Unit	$w_i w_{ij}^*$	Y_1	Y_2	$w_i^{(1)} w_{ij}^{*(1)}$	$w_i^{(2)} w_{ij}^{*(2)}$...	$w_i^{(5)} w_{ij}^{*(5)}$
1	0.10	2	1	0	0.111		0.111
2	0.10	2	2	0.111	0		0.111
3	0.10	1	3	0.111	0.111		0.111
4	0.10	4	4	0.111	0.111		0.111
5	0.10	2	5	0.111	0.111		0
6	0.05	3	1*	$0.111 (0.5 - \phi_1)$	0.055		$0.111 (0.5 + \phi_5)$
	0.05	3	5*	$0.111 (0.5 + \phi_1)$	0.055		$0.111 (0.5 - \phi_5)$
7	0.10	5	3	0.111	0.111		0.111
8	0.10	7	6	0.111	0.111		0.111
9	0.10	8	9	0.111	0.111		0.111
10	0.05	5*	3*	0.055	0.055		0.055
	0.05	7*	6*	0.056	0.056		0.056

Example (Continued)

Unit	$w_i w_{ij}^*$	Y_1	Y_2	$w_i^{(6)} w_{ij}^{*(6)}$	$w_i^{(7)} w_{ij}^{*(7)}$...	$w_i^{(10)} w_{ij}^{*(10)}$
1	0.10	2	1	0.111	0.111		0.111
2	0.10	2	2	0.111	0.111		0.111
3	0.10	1	3	0.111	0.111		0.111
4	0.10	4	4	0.111	0.111		0.111
5	0.10	2	5	0.111	0.111		0.111
6	0.05	3	1*	0	0.055		0.055
	0.05	3	5*	0	0.056		0.056
7	0.10	5	3	0.111	0		0.111
8	0.10	7	6	0.111	0.111		0.111
9	0.10	8	9	0.111	0.111		0.111
10	0.05	5*	3*	0.55	$0.111(0.5 - \phi_7)$		0
	0.05	7*	6*	0.56	$0.111(0.5 + \phi_7)$		0

Variance estimation for fractional imputation

- Variance estimator is a function of ϕ_k 's :

$$\hat{V}_\phi = \sum_{k=1}^L c_k \left(\sum_{i \in A_R} \alpha_{\phi, i}^{(k)} y_i - \sum_{i \in A_R} \alpha_i y_i \right)^2$$

- Naive variance estimator ($\phi_k \equiv 0$) : Underestimation
- Increasing the ϕ_k will increase the value of variance estimator
- How to decide ϕ_k ?

$$E \left(\hat{V}_\phi \right) - \text{Var} \left(\hat{\theta}_I \right) = E \left\{ \sum_{g=1}^G \sum_{i \in A_R} \left[\sum_{k=1}^L c_k \left(\alpha_{\phi, i}^{(k)} - \alpha_i \right)^2 - \alpha_i^2 \right] \sigma_g^2 \right\}$$

Variance estimation for fractional imputation

- Kim and Fuller (2004) showed that if

$$\sum_{i \in A_R} w_{ij}^{*(k)} = 1 \quad (\text{C.1})$$

and

$$\sum_{i \in A_{Rg}} \sum_{k=1}^L c_k \left(\alpha_i^{(k)} - \alpha_i \right)^2 = \sum_{i \in A_{Rg}} \alpha_i^2, \quad (\text{C.2})$$

then the replication variance estimator defined by

$$\hat{V}_I = \sum_{k=1}^L c_k \left(\hat{\theta}_I^{(k)} - \hat{\theta}_I \right)^2,$$

where $\hat{\theta}_I^{(k)} = \sum_{i \in A_R} \alpha_i^{(k)} y_i$, is unbiased for the total variance under the cell mean model.

Simulation study

Simulation Setup

- Three variables of size $n = 300$ are generated:

$$Y_1 \sim U(0, 2)$$

$$Y_2 = 1 + Y_1 + e_2$$

$$Y_3 = 2 + Y_1 + 0.5y_2 + 0.5e_3$$

where e_2 and e_3 are independently generated from a standard normal distribution truncated outside $[-3, 3]$.

- Response indicator functions for Y_1, Y_2, Y_3 :

$$\delta_1, \delta_2, \delta_3 \stackrel{i.i.d.}{\sim} \text{Bernoulli}(0.7)$$

- Interested in $\theta_1 = E(Y_1)$, $\theta_2 = E(Y_2)$, $\theta_3 = E(Y_3)$, $\theta_4 = P(Y_1 < 1, Y_2 < 2)$, and $\theta_5 = E(Y_3 \mid D = 1)$ with $D \sim \text{Bernoulli}(0.3)$.

Simulation study

Simulation Setup

- $B = 2,000$ simulation samples.
- Multivariate fractional hot deck imputation was used with $m = 5$ fractional imputation.
- Categorical transformation (with 4 categories) was used to each of Y_1 , Y_2 , and Y_3 .
- Within imputation cell, joint hot deck imputation was used.

Simulation study

Results: Point estimation

Table 1 Point estimation

Parameter	Method	Mean	Std Var.
θ_1	Complete Data	1.00	100
	FHDI	1.00	135
θ_2	Complete Data	2.00	100
	FHDI	2.00	142
θ_3	Complete Data	4.00	100
	FHDI	4.00	132
θ_4 Proportion	Complete Data	0.34	100
	FHDI	0.33	140
θ_5 Domain Mean	Complete Data	4.00	100
	FHDI	4.00	97

Simulation study

Results: Variance estimation

Table 2 Variance estimation

Parameter	Relative Bias (%)
$V(\hat{\theta}_1)$	6.69
$V(\hat{\theta}_2)$	-0.79
$V(\hat{\theta}_3)$	4.61
$V(\hat{\theta}_4)$	3.82
$V(\hat{\theta}_5)$	6.50

- Fractional hot deck imputation is considered.
- Categorical data transformation was used for approximation.
 - Does not rely on parametric model assumptions.
 - Two-stage imputation
 - 1 Stage 1: Imputation of cells (using conditional cell probability)
 - 2 Stage 2: Joint hot deck imputation within imputation cell.
- Replication-based approach for imputation variance estimation.
- Useful for general-purpose estimation, including domain estimation.
- To be implemented in SAS: **Proc SurveyImpute**