

# New Methodologies for the Study and Decomposition of Interviewer Effects in Surveys

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# Overview: Research Questions

- **RQ1:** What models can we use to study interviewer effects in the absence of interpenetrated designs?
  - Inflating variances for intra-interviewer correlations arising entirely from sample assignment leads to erroneous inferences!
- **RQ2:** If “total” interviewer variance arises from a combination of nonresponse error variance and measurement error variance among interviewers, what models can we use to study both error sources *simultaneously*?

# Review of the Literature

- Statistical Methods for estimating correlated components of variance in the presence of non-interpenetrated designs:
  - Biemer and Stokes (1985)
  - Kleffe et al. (1991)
  - Gao and Smith (1998)
  - Von Sanden and Steel (2008)
- These studies assume *semi*-interpenetrated designs
- Common practice involves fitting multilevel models including area-specific covariates (Schaeffer et al. 2010)

# Review of the Literature

- Decomposing total interviewer variance into measurement error variance and nonresponse error variance:
  - West and Olson (2010): telephone surveys
  - West, Kreuter, and Jaenichen (2013): FTF surveys
- These studies have followed a very simple descriptive approach using multilevel models, assuming interpenetration and independence of the two error sources within interviewers

# Gaps in the Literature

- Methods for the estimation of interviewer variance components are needed in designs where there is complete (or nearly complete) lack of interpenetration (most FTF surveys)
- More elegant modeling methods are also needed for studying the decomposition of total interviewer variance into the separate error variance components

# Method 1: Anchoring

- Consider a simple random sample...
- If cases with correlated values on a variable of interest are assigned to interviewers in a non-random fashion, we are just re-ordering the random sample given agents of the data collection process
- We have not altered anything about the actual data: no interviewer effects, no variance inflation
- Adjusting variance estimates for “interviewer” effects would lead to anti-conservative inferences

# Method 1: Anchoring (cont'd)

- **Basic idea:** Identify an ancillary variable (“anchor”) that
  1. Is unlikely to be subject to interviewer effects in measurement (e.g., age)
  2. Is correlated with a key survey variable of interest that may be subject to interviewer effects
- Next, fit a model allowing the two variables to have correlated residuals, and including random interviewer effects ONLY for the survey variable
- This removes the portion of the within-interviewer correlation due to non-random assignment, leaving a “clean” estimate of the between-interviewer variance



# Method 1: Anchoring (cont'd)

- In the simplest case, we have two variables, one ( $Y_1$ ) treated as measurement error-free, and one ( $Y_2$ ) treated as possibly having interviewer-induced measurement error
- We fit a model of the form

$$y_{ijk} = \mu_k + I(k = 2)b_i + \varepsilon_{ijk}$$

where  $i$  indexes interviewers,  $j$  indexes respondents within interviewers,  $k$  indexes the variable,  $b_i \sim N(0, \tau^2)$  and  $\begin{pmatrix} \varepsilon_{ij1} \\ \varepsilon_{ij2} \end{pmatrix} \sim N\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{bmatrix} \sigma_1^2 & \sigma_{12} \\ \sigma_{12} & \sigma_2^2 \end{bmatrix}\right)$ .

# Method 1: Anchoring (cont'd)

- Standard linear mixed model software can be used to obtain a REML point estimate of the mean for the second variable, along with an estimated variance component (we used PROC MIXED)
- High correlation between the residuals will lead to a more accurate estimate of the variance component, and thus of the true impact of the interviewer-induced measurement error on the variance of the estimated mean

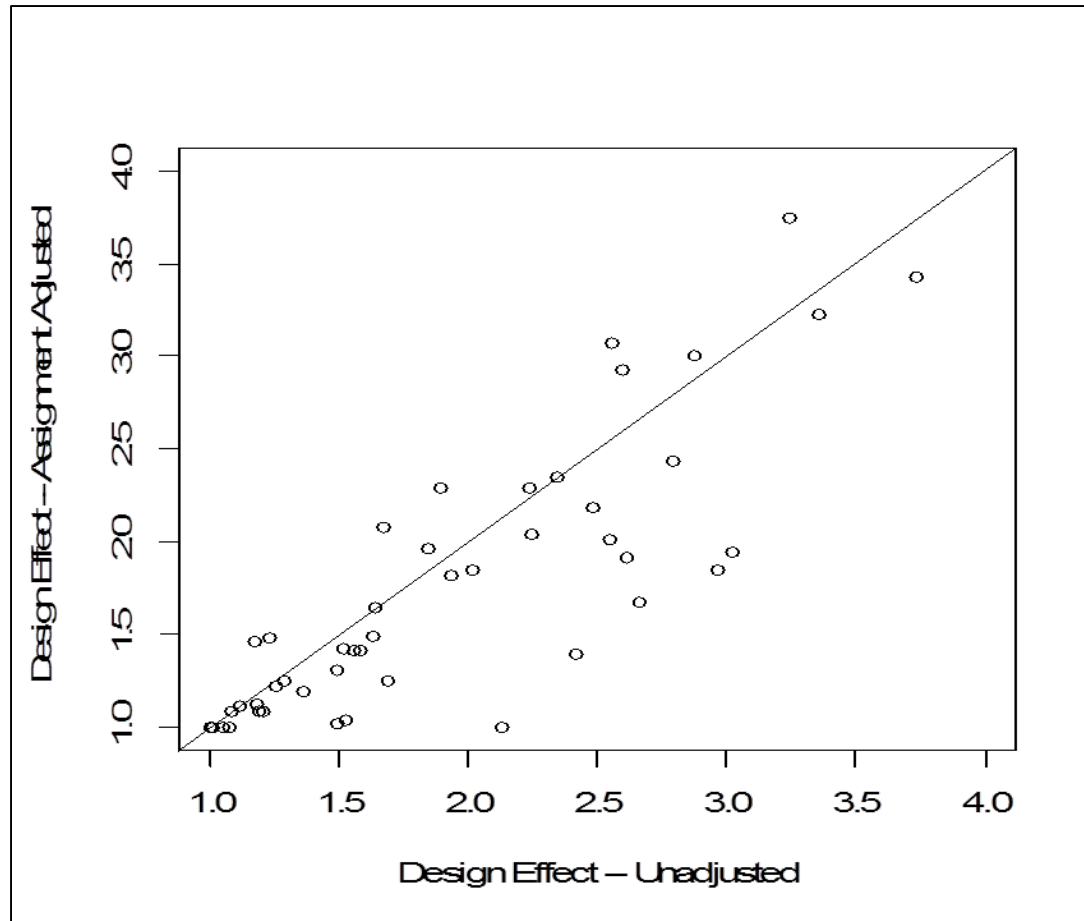
# Anchoring: An Illustration

- In a preliminary example using data from the 2012 BRFSS, we assume that the measurement error free variable is **age** and the variable of interest is **perceived health status** (1 = poor, ..., 5 = excellent)
- We choose age as an anchoring variable because:
  - a) We believe it is likely to be reported without differential measurement error,
  - b) It is associated with interviewer assignment, as interviewers tend to work shifts at different times of the day, and interview time of day is associated with age, and
  - c) It is also associated with health status.

# Anchoring: An Illustration (cont'd)

- We compute the interviewer effect for mean health status in each of the 50 states:
  - a) assuming an interpenetrated design, and
  - b) using the assignment-adjusted random effect, where age is assumed to be reported without measurement error
- The mean interviewer effect was 2.09 in the unadjusted analysis vs. 1.86 in the adjusted analysis
- Here is a picture of the results...

# Anchoring: An Illustration (cont'd)



- While most states were unchanged in their interviewer effect estimates, 9 states had decreases of more than 25%, including one of 53%!

# Multivariate Extensions

- For analysts interested in regression models, we can use standard software to fit the model

$$y_{ijk} = \beta_{0k} + \sum_{l=1}^p \beta_{lk} x_{ij} + b_{iK0} + \sum_{l=1}^p b_{iKl} x_{ij} + \varepsilon_{ijk},$$

with  $i = 1, \dots, I$ ,  $j = 1, \dots, J_i$ ,  $k = 1, \dots, K$ , where

$y_{ij1}, \dots, y_{ijK-1}$  are assumed free from measurement error,  $(b_{iK0} \cdots b_{iKp})^T \sim N_{p+1}(0, D)$ ,  $(\varepsilon_{ij1} \cdots \varepsilon_{ijK})^T \sim N_K(0, \Sigma)$ , and the variance-covariance matrices are unstructured.

# Method 2: Assignment Propensity

- An alternative possible approach to dealing with this problem adapts propensity score adjustment methods to develop “assignment weights,” equal to the inverse of the probability of assignment to a given interviewer
- The assignment probability is estimated as a function of covariates known to be (or treated as) free of measurement error

# Method 2: Assignment Propensity

- Use of these weights in estimation “re-creates” an approximate interpenetrated design
- The approximation will improve to the degree that the covariates capture all of the assignment mechanism that is correlated with the measurement error
- The weights would be estimated by fitting a multinomial model to respondent data, with interviewer ID as the dependent variable



# Method 2: Assignment Propensity

- If the covariates are unrelated to assignment, then the weights will be approximately equal (i.e., the design is essentially interpenetrated)
- If there are strong associations of the covariates with assignment, the weights should remove the impact of the assignment from analysis
- If sampling weights are present, these would be used for estimation of the multinomial model, and the sampling weights would be multiplied by the assignment weights to obtain final analysis weights

# Method 3: MLMI

- MLMI = A Multi-Level Multiple Imputation approach to decomposing total interviewer variance
- We need a *simultaneous* modeling method for decomposing total interviewer variance into measurement error and nonresponse error variance, while also enabling estimation of the covariance of these two error sources (among interviewers)

# Method 3: MLMI

- Outline of the proposed approach:
  1. Form a data set using the full sample, including true values of  $Y$  (for the full sample) and reported values of  $Y$  for respondents (note: generally **rare** to have true  $Y$  values for the full sample)
  2. Create a (1, -1) variable ( $X$ ), where 1 = respondent, and -1 = non-respondent. Ignoring interviewer effects, fit a regression model to the true values of  $Y$  using the (1, -1) variable, which will produce the full sample mean ( $B_0$ ) and the nonresponse error ( $B_1$ )

# Method 3: MLMI

3. Allowing  $B_0$  to randomly vary across interviewers will capture variance in assignment (should be zero for interpenetrated assignment); allowing  $B_1$  to vary across interviewers will capture variance in the nonresponse errors; we can also allow the random effects to covary!
4. Non-respondents will have missing values on *reported Y*; use true values, interviewer effects, and other auxiliary variables to *impute* reported values of  $Y$  for non-respondents

# Method 3: MLMI

5. In each imputed data set, fit a model to the *reported values* of Y (including the imputed values), allowing the intercept and the “response” effect to randomly vary across interviewers
6. **Assuming interpenetration**, we can now estimate the variance of the intercepts (measurement error), the variance of the “response” effects (nonresponse error), and the covariance of these effects across interviewers
  - NOTE: we could consider the assignment propensity approach if interpenetration was not evident...

# Method 3: MLMI

- Clearly the success of this method depends on some key features of the available data:
  - A rich sampling frame including true values on selected variables
  - Auxiliary variables that are strongly predictive of survey reports (for good imputation models)
- We are eager to apply this approach to existing data sets (which again will be rare)...this is the next (methodological) step

# Discussion Points

- What does everyone think about these ideas?
- How would you suggest that we refine these approaches?
- Are we missing any other developments in the literature in these areas?
- Empirical applications are certainly needed!  
We are just at the idea phase right now.